SUPPLEMENTAL DIGITAL CONTENT - TECHNICAL APPENDICES

Table of Contents for the Technical Appendix

Appendix A: Narrative Description of Model Construction

Appendix B: Appendix B: CD4 state transitions parameters and CD4 slope

Appendix C: Pre-ART Mortality Parameters

Appendix D: On-ART Mortality Parameters

Appendix E: Whole cohort analysis with lognormal distribution around ART delay

Appendix A: Narrative Description of Model Construction

The following text describes, in greater detail, the steps taken in developing the state-transition model. First, we developed a theoretical model, defining the transition states necessary to generate estimates of HIV mortality as a function of ART delay (Figure 1, main manuscript text). Next, we used patient-level data from three South African cohorts to calculate transition probabilities (and associated variability estimates) of transitions between CD4 count states with an *a priori* width of 50 cells/mm³, starting at 350. We also estimated the within-stratum mortality risk according to CD4 cell state; this process required differentiation between pre-ART mortality and on-ART mortality.

Model Development

Once all state transition probabilities (and estimates of variability) were defined, we used TreeAge 2011 to build a Markov state transition model with a cycle duration of 1 week allowing for transitions from the pre-ART to on-ART and between CD4 states during the pre-ART period. In our primary analysis, we estimated mortality for each of our pre-specified discrete delays (e.g., no delay, 6 weeks, 12 weeks, etc.) within each of our CD4 count states (e.g., 0-24, 25-49, 50-99, etc.), in order to provide a patient-level estimate of the impact of various delays according to entry CD4 count. We also considered a "whole-cohort" analysis, in which we took a population that reflected the distribution of our pre-ART cohort and assigned each individual to a delay using random draws from a log-normal distribution, centered at our pre-specified delay levels (e.g., 6 weeks, 12 weeks, etc.). This analysis was designed to evaluate the population-level effects of reducing delay, accounting for the right-skew in ART delay that may be more reflective of a "real-world" population. For each scenario based on timing in ART initiation and CD4 count and clinic entry we performed a micro-simulation of 1000 patients in 10,000 trials to achieve convergence of results. We used the mean and upper 97.5 and lower 2.5 range to describe the 95% confidence interval.

CD4 count state transition probabilities

Description of cohort from which state transition probabilities were calculated The cohort was created using patient level data from a clinical trial of isoniazid preventive therapy among gold miners at a single mine site in South Africa. Employees with a prior HIV-positive test were recruited to participate from a workplace wellness clinic. None of the participants received antiretroviral therapy (ART) during the course of the trial because ART was not available at that time in South Africa. 1863 participants were included with a median follow-up time of 1 year (IQR: 0.7, 2.1). CD4 count was enumerated a median of every 6 months. Total mortality was 39/1863 (2%).

Calculation of transition probabilities in the reference case

We calculated transition probabilities between CD4 count states (<25, 25-49, 50-99, 100-149, 150-199, 200-249, 250-299, 300-349, >350) by calculating the Nelson-Aalen hazard function defining a "failure" as

progression to a lower state and separately defining "failure" as a transition to a higher state. We used the four week hazard function and divided by four to arrive at the one week parameter.

Calculation of variability

Since a simple binomial distribution would underestimate the true variability in CD4 transition probabilities across individuals, we estimated the variability in these transition parameters by first performing a linear regression of the instantaneous hazard of transition against the midpoint CD4 cell count in each stratum, excluding the >350 cells/mm³ stratum and weighted by the number of observations in each stratum. This regression generated an equation for the decline in CD4 count as a function of CD4 count. Second, we constructed a symmetric beta distribution (shape parameter [alpha] pre-specified at 4) around the slope of this regressed line (i.e., predicted change in CD4 count per week, as a function of existing CD4 count), with an upper bound that would provide the probabilities of upward transition seen in our actual data. We chose a beta distribution, rather than a normal distribution, to provide replicability of our transition probabilities while avoiding the potential for unrealistic scenarios (e.g., a CD4 decline of 60 cells/week in a patient with a CD4 count of 55). Third, we took random draws from this beta distribution using Monte Carlo simulation and estimated the probability of transition as (change in CD4 count)/(width of CD4 interval). Thus, a patient with a weekto-week change in CD4 count of 3 cells/mm³ in a stratum of width 50 cells/mm³ (e.g., 201-250 cells/mm³) would have a probability of downward transition of (3/50) = 6%. By simulating thousands of patients until results converged, we were able to estimate the degree of variability in our transition probabilities.

Comparison with the literature

Using our equation for the CD4 count slope, for a CD4 count of 75, the expected slope is -71 cells/mm³/year. We compared this slope to values reported in the literature (Table B-2 in the Appendix).

Weekly probability of death, pre-ART

Description of cohort from which pre-ART mortality probabilities were calculated

We used patient level data from a community-based HIV management program operated through freeof-charge care provided by general practitioners in three provinces of South Africa. Inclusion criteria were ART-naïve and known date of first CD4 count enumeration in the clinic. We supplemented data from this cohort with a linkage to the South African Department of Home Affairs vital statistics registry to identify unreported deaths. We had national ID numbers for 65% patients allowing for linkage. To reduce bias and underestimated of mortality, we used inverse probability weighting to estimate mortality for those patients who were both lost from care and lacked a valid ID number.

Calculation of transition (mortality) probabilities in the reference case

We used survival time analysis with entry into observation on the date of the first CD4 assay and exit at the date of death with censoring one month prior to the linkage with the vital statistics register or ART initiation. As we were able to ascertain our outcome of interest, death, through means beyond clinic

follow-up, we did not censor patients who were lost from care. We calculated the Nelson-Aalen mortality hazard function by CD4 state using time-updated CD4 count values for the six month period and divided by 26 to obtain a weekly hazard of mortality. We did this after evaluating the mortality hazard by 4 week intervals and not observing a difference over the six month period.

Calculation of variability

We used a beta distribution around the calculated weekly mortality hazard and the standard deviation of the mortality hazard for probabilistic sampling using Monte Carlo simulations.

Comparison with the literature

We simulated mortality, by CD4 state, for 52 weeks without ART to provide pre-ART mortality. We converted raw mortality by week to mortality by 100 person-years for easier comparability with the published literature (Table 2, manuscript).

Weekly probability of death, after initiating ART

Description of cohort from which on-ART mortality probabilities were calculated

We used patient level data from a community-based HIV management program operated through freeof-charge care provided by general practitioners in three provinces of South Africa and a multi-site workplace program. Both programs used similar management algorithms and shared common data collection forms and a relational database. Inclusion criteria were ART-naïve, WHO clinic stage <4, and age ≥18 years old. We supplemented data from this cohort with a linkage to the South African Department of Home Affairs vital statistics registry to identify unreported deaths. We had national ID numbers for 80% of patients allowing for linkage. To reduce bias and underestimation of mortality, we used inverse probability weighting to estimate mortality for those patients who were both lost from care and lacked a valid ID number.

Calculation of mortality probabilities

We used survival time analysis with entry into observation on the date of ART initiation and exit at the date of death with censoring one month prior to vital statistics linkage closure. We calculated the Nelson-Aalen mortality hazard function by CD4 state using time-updated CD4 count values. We used four week probabilities to generate an exponential equation to describe the mortality hazard from the time of ART initiation to 52 weeks on ART. We then calculated the mortality for the midpoint of each one-week interval to generate mortality probabilities.

Calculation of variability in mortality probability

We used a Normal distribution around the calculated weekly mortality hazard and the standard deviation of the mortality hazard for probabilistic sampling during Monte Carlo simulations.

Comparison with the literature

We simulated mortality, by CD4 state, for 52 weeks on ART to provide on-ART mortality. We converted raw mortality by week to mortality by 100 person-years for easier comparability with the published literature (Table 2, manuscript).

Appendix B: CD4 state transitions parameters and CD4 slope

| Starting CD4 | Probability of transition per week | | | | | | |
|--------------|------------------------------------|-----------|-------------|--|--|--|--|
| state | | | | | | | |
| | Down 1 state* | No change | Up 1 state* | | | | |
| <25 | 0 | 0.998 | 0.0012 | | | | |
| 25-49 | 0.0447 | 0.941 | 0.0136 | | | | |
| 50-99 | 0.0203 | 0.972 | 0.00667 | | | | |
| 100-149 | 0.0202 | 0.973 | 0.00662 | | | | |
| 150-199 | 0.0189 | 0.974 | 0.00675 | | | | |
| 200-249 | 0.0171 | 0.976 | 0.00631 | | | | |
| 250-299 | 0.0145 | 0.979 | 0.00596 | | | | |
| 300-349 | 0.0152 | 0.978 | 0.00612 | | | | |
| ≥350 | 0.0099 | 0.99 | 0 | | | | |

Table B1: CD4 transition probabilities

* The mean of the standard deviations for each of the downward transition probabilities was: 0.0055 and for the upward probabilities was 0.0013.

Linear regression CD4 count slope equation: CD4 count slope: y = 1.486 – 0.0016 * CD4 count + e

| STUDY | POPULATION | Modifying parameter | CD4 decline, cells/mm ³ /yr (95% Cl) |
|--------------------------------------|---|-------------------------------|--|
| This study | South African Miners clinical trial of TB prevention | CD4: 75 | -71 (-79, -63) |
| Mellors et al Ann Int Med 1997 | US MACS | VL: 10,000-30,000 | -64 (-70.0, -59.6) |
| | | VL: >30,000 | -76.5 (-70.5, -82.9) |
| Kiwanuka et al. JAIDS 2010 | Uganda | | -34.5 (-47, -22) |
| Mussini et al. AIDS 2011 | US Cohort | 2-8 years from ART initiation | -39.3 (-44.3, -34.3) |
| | | <2 years from ART initiation | -96.6 (-100.9, -92.6) |
| Holmes et al. JAIDS 2006 | South Africa Clinical cohort (10% received AZT and 31% died during f/u) | CD4 >500 | -47.1 (-54, -40) |
| | | 351-500 | -30.6 (-37, -23) |
| | | 201-350 | -20.5 (-13.7, -27) |
| Ding et al. JAIDS 2009 | India, Clinical cohort | | -97.1 (+/- 11.4) |

Table B-2: Comparison of studies of CD4 slope among ART-naïve individuals

Appendix C: Pre-ART Mortality Parameters

| | <25 | 25-49 | 50-99 | 100-149 | 150-199 | 200-249 | 250-299 | 300-349 | ≥350 |
|--------------|-------|-------|---------|---------|---------|---------|---------|----------|----------|
| Mortality | 0.014 | 0.01 | 0.00482 | 0.00346 | 0.00249 | 0.00179 | 0.00128 | 0.000924 | 0.000664 |
| Probability, | | | | | | | | | |
| per week | | | | | | | | | |
| Standard | 0.002 | 0.002 | 0.002 | 0.001 | 0.001 | 0.001 | 0.001 | 0.001 | 0.0001 |
| Deviation | | | | | | | | | |

Table C1: Pre-ART Mortality Transition Probabilities by current CD4 count

Appendix D: On-ART Mortality Parameters

| Table D1: On-ART Mortality Transition Probabilities by weeks from ART initiation and CD4 count at |
|---|
| ART initiation |

| Weeks | <25 | 25-49 | 50-99 | 100-149 | 150-199 | 200-249 | 250-299 | 300-349 | ≥350 |
|-----------------|--------|--------|--------|---------|---------|---------|---------|---------|--------|
| from | | | | | | | | | |
| ART | | | | | | | | | |
| initia- tion | | | | | | | | | |
| 1 | 0.0101 | 0.0067 | 0.0049 | 0.0035 | 0.0025 | 0.0017 | .0013 | 0.0009 | 0.0007 |
| 2 | 0.0098 | 0.0066 | 0.0047 | 0.0034 | 0.0024 | 0.0017 | 0.0012 | 0.0009 | 0.0007 |
| - | 0.0095 | 0.0064 | 0.0046 | 0.0033 | 0.0024 | 0.0016 | 0.0012 | 0.0008 | 0.0007 |
| 4 | 0.0092 | 0.0062 | 0.0045 | 0.0032 | 0.0023 | 0.0016 | 0.0012 | 0.0008 | 0.0007 |
| 5 | 0.0089 | 0.0061 | 0.0044 | 0.0031 | 0.0022 | 0.0016 | 0.0011 | 0.0008 | 0.0006 |
| 6 | 0.0087 | 0.0059 | 0.0043 | 0.0030 | 0.0022 | 0.0015 | 0.0011 | 0.0008 | 0.0006 |
| 7 | 0.0084 | 0.0058 | 0.0041 | 0.0029 | 0.0021 | 0.0015 | 0.0011 | 0.0008 | 0.0006 |
| 8 | 0.0082 | 0.0056 | 0.0040 | 0.0029 | 0.0020 | 0.0014 | 0.0010 | 0.0008 | 0.0006 |
| 9 | 0.0080 | 0.0055 | 0.0039 | 0.0028 | 0.0020 | 0.0014 | 0.0010 | 0.0007 | 0.0006 |
| 10 | 0.0077 | 0.0053 | 0.0038 | 0.0027 | 0.0019 | 0.0013 | 0.0010 | 0.0007 | 0.0006 |
| 11 | 0.0075 | 0.0052 | 0.0037 | 0.0026 | 0.0019 | 0.0013 | 0.0009 | 0.0007 | 0.0006 |
| 12 | 0.0073 | 0.0051 | 0.0036 | 0.0025 | 0.0018 | 0.0013 | 0.0009 | 0.0007 | 0.0006 |
| 13 | 0.0071 | 0.0049 | 0.0035 | 0.0025 | 0.0017 | 0.0012 | 0.0009 | 0.0007 | 0.0006 |
| 14 | 0.0069 | 0.0048 | 0.0034 | 0.0024 | 0.0017 | 0.0012 | 0.0009 | 0.0007 | 0.0006 |
| 15 | 0.0067 | 0.0047 | 0.0033 | 0.0023 | 0.0016 | 0.0011 | 0.0008 | 0.0006 | 0.0006 |
| 16 | 0.0065 | 0.0046 | 0.0033 | 0.0023 | 0.0016 | 0.0011 | 0.0008 | 0.0006 | 0.0005 |
| 17 | 0.0063 | 0.0044 | 0.0032 | 0.0022 | 0.0015 | 0.0011 | 0.0008 | 0.0006 | 0.0005 |
| 18 | 0.0061 | 0.0043 | 0.0031 | 0.0021 | 0.0015 | 0.0010 | 0.0008 | 0.0006 | 0.0005 |
| 19 | 0.0060 | 0.0042 | 0.0030 | 0.0021 | 0.0014 | 0.0010 | 0.0007 | 0.0006 | 0.0005 |
| 20 | 0.0058 | 0.0041 | 0.0029 | 0.0020 | 0.0014 | 0.0010 | 0.0007 | 0.0006 | 0.0005 |
| 21 | 0.0056 | 0.0040 | 0.0028 | 0.0020 | 0.0014 | 0.0010 | 0.0007 | 0.0006 | 0.0005 |
| 22 | 0.0055 | 0.0039 | 0.0028 | 0.0019 | 0.0013 | 0.0009 | 0.0007 | 0.0006 | 0.0005 |
| 23 | 0.0053 | 0.0038 | 0.0027 | 0.0018 | 0.0013 | 0.0009 | 0.0007 | 0.0005 | 0.0005 |
| 24 | 0.0051 | 0.0037 | 0.0026 | 0.0018 | 0.0012 | 0.0009 | 0.0006 | 0.0005 | 0.0005 |
| 25 | 0.0050 | 0.0036 | 0.0025 | 0.0017 | 0.0012 | 0.0009 | 0.0006 | 0.0005 | 0.0005 |
| 26 | 0.0049 | 0.0035 | 0.0025 | 0.0017 | 0.0012 | 0.0008 | 0.0006 | 0.0005 | 0.0005 |
| 27 | 0.0047 | 0.0034 | 0.0024 | 0.0016 | 0.0011 | 0.0008 | 0.0006 | 0.0005 | 0.0005 |
| 28 | 0.0046 | 0.0033 | 0.0024 | 0.0016 | 0.0011 | 0.0008 | 0.0006 | 0.0005 | 0.0004 |
| 29 | 0.0045 | 0.0033 | 0.0023 | 0.0016 | 0.0011 | 0.0008 | 0.0006 | 0.0005 | 0.0004 |
| 30 | 0.0043 | 0.0032 | 0.0022 | 0.0015 | 0.0010 | 0.0007 | 0.0005 | 0.0005 | 0.0004 |
| 31 | 0.0042 | 0.0031 | 0.0022 | 0.0015 | 0.0010 | 0.0007 | 0.0005 | 0.0005 | 0.0004 |

| 32 | 0.0041 | 0.0030 | 0.0021 | 0.0014 | 0.0010 | 0.0007 | 0.0005 | 0.0004 | 0.0004 |
|----|--------|--------|--------|--------|--------|--------|--------|--------|--------|
| 33 | 0.0040 | 0.0029 | 0.0021 | 0.0014 | 0.0009 | 0.0007 | 0.0005 | 0.0004 | 0.0004 |
| 34 | 0.0038 | 0.0029 | 0.0020 | 0.0013 | 0.0009 | 0.0006 | 0.0005 | 0.0004 | 0.0004 |
| 35 | 0.0037 | 0.0028 | 0.0019 | 0.0013 | 0.0009 | 0.0006 | 0.0005 | 0.0004 | 0.0004 |
| 36 | 0.0036 | 0.0027 | 0.0019 | 0.0013 | 0.0009 | 0.0006 | 0.0005 | 0.0004 | 0.0004 |
| 37 | 0.0035 | 0.0026 | 0.0018 | 0.0012 | 0.0008 | 0.0006 | 0.0004 | 0.0004 | 0.0004 |
| 38 | 0.0034 | 0.0026 | 0.0018 | 0.0012 | 0.0008 | 0.0006 | 0.0004 | 0.0004 | 0.0004 |
| 39 | 0.0033 | 0.0025 | 0.0017 | 0.0012 | 0.0008 | 0.0006 | 0.0004 | 0.0004 | 0.0004 |
| 40 | 0.0032 | 0.0024 | 0.0017 | 0.0011 | 0.0008 | 0.0005 | 0.0004 | 0.0004 | 0.0004 |
| 41 | 0.0031 | 0.0024 | 0.0017 | 0.0011 | 0.0007 | 0.0005 | 0.0004 | 0.0004 | 0.0004 |
| 42 | 0.0031 | 0.0023 | 0.0016 | 0.0011 | 0.0007 | 0.0005 | 0.0004 | 0.0004 | 0.0004 |
| 43 | 0.0030 | 0.0023 | 0.0016 | 0.0010 | 0.0007 | 0.0005 | 0.0004 | 0.0003 | 0.0004 |
| 44 | 0.0029 | 0.0022 | 0.0015 | 0.0010 | 0.0007 | 0.0005 | 0.0004 | 0.0003 | 0.0003 |
| 45 | 0.0028 | 0.0021 | 0.0015 | 0.0010 | 0.0006 | 0.0005 | 0.0004 | 0.0003 | 0.0003 |
| 46 | 0.0027 | 0.0021 | 0.0014 | 0.0009 | 0.0006 | 0.0005 | 0.0003 | 0.0003 | 0.0003 |
| 47 | 0.0026 | 0.0020 | 0.0014 | 0.0009 | 0.0006 | 0.0004 | 0.0003 | 0.0003 | 0.0003 |
| 48 | 0.0026 | 0.0020 | 0.0014 | 0.0009 | 0.0006 | 0.0004 | 0.0003 | 0.0003 | 0.0003 |
| 49 | 0.0025 | 0.0019 | 0.0013 | 0.0009 | 0.0006 | 0.0004 | 0.0003 | 0.0003 | 0.0003 |
| 50 | 0.0024 | 0.0019 | 0.0013 | 0.0008 | 0.0006 | 0.0004 | 0.0003 | 0.0003 | 0.0003 |
| 51 | 0.0023 | 0.0018 | 0.0013 | 0.0008 | 0.0005 | 0.0004 | 0.0003 | 0.0003 | 0.0003 |
| 52 | 0.0023 | 0.0018 | 0.0012 | 0.0008 | 0.0005 | 0.0004 | 0.0003 | 0.0003 | 0.0003 |

Appendix E: Whole cohort analysis with lognormal distribution around ART delay

The table (Table E1) represents results of whole cohort simulations when using a distribution around the delay duration. We used a lognormal distribution with the mean the "delay" and the median 25% lower than the mean. This distribution provides a long tail of delay in ART initiation.

| Weeks to ART | 0 | 6 | 10 | 26 | 52 | | | | |
|-------------------------|--------------|-----------------|-----------------|---------------|--------------|--|--|--|--|
| Total mortality | 11 (7.4, 15) | 13 (8.8, 17) | 14 (10. 18) | 18 (14, 21) | 22 (19, 26) | | | | |
| Excess mortality | 0 | 1.5 (-4.1, 7.1) | 2.6 (-2.9, 8.1) | 6.5 (1.3, 12) | 11 (5.7, 16) | | | | |

Table E1: Mortality by ART delay using a beta distribution around the duration of the delay